# PORTFOLIO DIVERSIFICATION OF NYSE

# SPECIALIST UNITS

bу

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The contents of this paper are solely the responsibility of the author.

Until recently each stock on the NEW-YORK Stock Exchange (NYSE) has been assigned to a specific specialist unit that has the responsability of dealing for its own account so as to maintain a "fair and orderly market". A specialist unit is a joint operation of one or more individual specialists. As of May 31, 1975 they were 381 individual specialists aggregated into 67 specialist units. Although competition among these dealer firms has recently been permitted, the fact remains that most specialist units have a monopoly, at least on the floor of the NYSE.

A system of monopoly dealers tends, for a given number of specialist units, to minimize the number of stocks per unit and therefore limits the diversification against risk that each unit can achieve. The less diversified the unit, the more volatile bid and ask prices will be relative to the "true" price of the stock because the total price adjustment to an inventory change would tend to be in the stock traded and could not be passed on to as many other stocks in the form of changes in bid-ask quotations for those stocks. If the unit specializes in a single stock, it can eliminate the inventory risk of holding that stock only by selling that stock. If the unit has a portfolio of stocks, it can eliminate the inventory risk of holding a stock by selling that stock or by selling other specialty stocks the returns of which are correlated with the return of the stock in inventory. Thus, in the second case, bid and ask prices change in all stocks whose returns are correlated with those of the stock traded in order to encourage transactions in those stocks as well. And the bid and ask price changes are less for the traded stock than they would be if the unit specialized in only that stock.

The purpose of this paper is to report on the degree of diversification of the 67 specialist units existing at the end of 1974. In the next section the data are described. Then the procedures for measuring diversification are presented. Finally the results and conclusions are given.

## DATA

The specialist unit assigned to each security is identified in the NYSE publication, Stocks and Specialists. The data for December 31, 1974 are utilized, and the analysis is restricted to common stocks only. Thus the specialist unit in preferred stocks is omitted from consideration. The specialist units analysed are listed in Table 1.

Measures of portfolio risk and return are based on measures of risk and return for the individual speciality stocks of each unit. These measures of risk and return are calculated using daily returns for the two year period, January 1, 1973 to December 31, 1974. Adjustment is made for all splits and stock dividends and account is taken of all cash dividend payments. The market index used in the calculation of risk measures is a value weighted index of all stocks.

# MEASURING DIVERSIFICATION

The measurement of diversification is based on the market model, which posits that the relation between the return on an individual stock,  $R_{i}$ , and the return on the market portfolio,  $R_{m}$  is:

$$R_{i} = a_{i} + b_{i} R_{m} + e_{i}$$
 (1)

where a, and b, are constants specific to the stock and where:

$$E(e_{i}) = 0, cov(R_{m}, e_{i}) = 0.$$

With these assumptions, one can write:

$$\sigma^{2}(R_{i}) = b_{i}^{2} \sigma^{2}(R_{m}) + \sigma^{2}(e_{i}),$$
 (2)

where  $b_i^2 \sigma^2(R_m)$  is usually referred to as "systematic risk" and  $\sigma^2(e_i)$  as "unsystematic risk". The second term is referred to as unsystematic risk because it is assumed that  $cov(e_i, e_j) = 0$  and that there is only one source of systematic variability – the behavior of the market as a whole.

The return on the portfolio of specialist unit k is given as:

$$R_{k} = \sum_{i=1}^{n} X_{i,k} R_{i,k} = \sum_{i=1}^{n} X_{i,k} (a_{i} + b_{i}R_{m} + e_{i})$$

$$R_{k} = \sum_{i=1}^{n_{k}} X_{i,k} a_{i} + \sum_{i=1}^{n_{k}} X_{i,k} b_{i,m} + \sum_{i=1}^{n_{k}} X_{i,k} e_{i}$$
(3)

and

$$\sigma^{2}(R_{k}) = b_{k}^{2} \sigma^{2}(R_{m}) + \sigma^{2}(e_{k})$$
 (4)

where

 $n_{i}$  = number of stocks in portfolio of dealer k.

X; = proportion of portfolio invested in stock i.

$$b_{k} = \sum_{i=1}^{n_{k}} X_{i,k} b_{i}$$

$$\sigma^{2}(e_{k}) = \sum_{i=1}^{n_{k}} X_{i}^{2} \sigma^{2}(e_{i})$$

The benefit of diversification is that the term,  $\sigma^2(e_k)$  in (4) becomes very small as the number of stocks increases.

The primary measure of specialist diversification used is the percentage of told risk which is not diversified away:

$$D_{k} = \frac{\sigma^{2}(e_{k})}{\sigma^{2}(R_{k})}$$

Of course, this measure depends on the weights,  $X_i$ , assigned to each stock in the portfolio. Two weighting schemes are utilized: equal weights in which case  $\frac{1}{n_k}$  of each speciality stock is held, and proportional weights in which case speciality stocks are held in the same proportions as they exist in the market as a whole. Under proportional weighting  $X_i = \frac{V_i}{T_i}$  when  $V_i = \max_{i=1}^{N} V_i$ 

value of common stock of firm i and T is the total number of firms with common stock listed on the NYSE.

These weighting schemes are essentially arbitrary since specialist units need not hold stocks in any particular proportion. In fact, specialists may have short positions in some stocks in order to hedge their portfolio.

Nevertheless, risk reduction will be easier the more opportunity the specialist unit has to diversify and measurement of differences among specialist units in their ability to diversify is the objective of this paper.

The failure of certain specialist units to be able fully to diversify need not, however, necessarily imply that NYSE rules against competing specialists limited diversification that would otherwise have taken place in a free market. Diseconomies of scale in trading many stocks could in principle offset the benefits of diversification. Thus even without portfolio restrictions dealers might choose non-diversified portfolios. Nevertheless the question still remains of whether the particular stocks chosen yield the maximum diversification. In order to answer this question, specialist portfolios are compared with randomly selected portfolios containing the same number of stocks. The stocks are selected with equal probability from all NYSE stocks. The same portfolio characteristics are calculated as for the portfolio based on the actual speciality stocks. The random selection procedure is replicated 25 times for each specialist unit, and the average values of the portfolio characteristics are reported.

## ESTIMATION OF THE MARKET MODEL

The use of daily data in estimating (1) directly can cause biases because of non-trading of securities. The less frequently a security is traded the more likely is the possibility that its return is measured over a time interval preceding that of the average time interval over which the market return is measured. This is due to the fact that returns are based on the last transaction price during the day, the time of which can vary across securities. 7

Scholes and Williams [.5] have devised a procedure of obtaining an unbiased, consistent estimator of  $b_i$  in (1), and their procedure was utilized in this study. This procedure requires calculation of slope coefficients for lagged relationship between  $R_i$  and  $R_m$ , and then averages the resulting values. Let

$$b_{i1} = \frac{\text{cov}(R_{it}, R_{mt})}{\sigma^2(R_{mt})}, b_{i2} = \frac{\text{cov}(R_{it}, R_{mt-1})}{\sigma^2(R_{mt-1})}$$

$$b_{i3} = \frac{\text{cov}(R_{it}, R_{mt+1})}{\sigma^2(R_{mt+1})}, b_4 = \frac{\text{cov}(R_{mt}, R_{mt-1})}{\sigma^2(R_{mt-1})}$$

Then the unbiased b used in this analysis is:

$$b_i = \frac{b_{i1} + b_{i2} + b_{i3}}{1 + 2 b_4}$$
,

and:
$$a_{i} = R_{i} - b_{i}R_{m}$$
 $e_{ir} = R_{ir} - a_{i} - b_{i}R_{m}$ 

This is equivalent to using an instrumental variable for R in (1) which is the sum of R and R m,t-1 and R m,t+1.

The results do not, however, appear to be substantially different than if (1) were estimated directly.

#### RESULTS

Portfolio characteristics for the 67 specialist units are given in Table 1 along with some data for randomly selected portfolios. The number of speciality stocks ranges from a low of 4 to a high of 93. The Beta coefficient,  $b_k$ , of specialist portfolio ranges from .727 to 1.174 for equal weighting and from .636 to 1.544 when market value weights are used. The column labeled  $D_{k}$ measures the proportion of told risk which could be diversified away with sufficient numbers of stocks, but was not. The column labeled  $\sigma^2(R_m)$  x 10,000 is the variance of the daily returns of the portfolio multiplied by 10,000. And the column labeled  $\overline{R}_{L}$  x 10,000 is the average daily rate of return to the portfolio during the two year period multiplied by 10,000. Returns are generally negative because the market was falling during the period analysed. Values of  $b_k$  and  $\overline{R}_k$  are not given for the randomly selected portfolios since the random selection procedure yields Beta values such that  $E(b_k) = 1$  and  $E(R_{t}) = E(R_{m})$ . Differences in "random"  $b_{t}$  across specialist units are due only to random errors, which are quite small because of the replication procedure.

Selected rank values on  $D_k$ , the proportion of total risk that was diversifiable are reported in Table 2 (in percentage terms). The table shows that, even when specialists hold all their stocks, a substantial number of units incurred risk that could have been diversified away with a portfolio containing more stocks. If stocks are held in equal proportions, 33 units incurred risk 24.01 % or more of which could have been diversified away. If stocks are held in market value weights, 33 units incurred risk 32.19 % or more of which could have been diversified away. The specialist unit with the most stocks (93) had the lowest level of diversifiable risk; the unit with the fewest stocks (4), the highest. Although these results indicate that specialists could have incurred less risk through greater diversification,

they do not show that it would have been optimal to do so since there may be offsetting costs in making markets in a large number of stocks. However, under competition, dealers would presumably be free to choose that number of specialty stocks which minimizes the cost of risk and their other costs, and this may be a number quite different than under the NYSE specialist system.

Figures 1 and 2 plot  $D_k$  against the number of stocks for equal weighted portfolios and for market weighted portfolios respectively. Each figure locates the specialist unit, denoted by S, and the random portfolio with the same number of stocks, denoted by R. An "X" indicates that two or more portfolios fall on the same point. First, the figure shows the benefits of diversification in the usual way — as the number of stocks increases diversifiable risk decreases. The cluster of random portfolios is "tighter" because of the replication procedure.

Second the figures indicate the extent to which specialist units incurred diversifiable risk greater or less than the amount one would expect on the basis of the number of stocks in the portfolio. In other words, suppose the specialist units were constrained only in the number of stocks. Do the stocks actually assigned to specialist units result in less diversifiable risk or is there little difference? The figures indicate that while some specialist units incur more diversifiable risk than the corresponding random portfolios, others incur less. The observed variation is consistent with random selection of specialist portfolios, and there is no systematic tendency for specialist portfolios to be more or less diversified than one would expect given the number of stocks in each portfolio. This is evident from the summary statistics in Table 1; the difference in mean  $\mathbf{D}_{\mathbf{k}}$  between specialist and random portfolios is not statistically significant.

#### SUMMARY

This existence of incomplete diversification for a significant number of specialist units is demonstrated on the basis of the market model. This result does not necessarily imply that complete diversification of specialist portfolios would be optimal since the benefits of diversification may be offset by diseconomies of scale arising from clerical and other costs of handling many stocks.

Second, supposing that limited diversification is desirable for these other reasons, the paper shows that the particular stocks chosen by specialist units provide no greater degree of diversification than if the same number of stocks had been chosen at random.

The effect of incomplete diversification on the price of dealer services is not investigated, although explicit measures of dealer diversification are provided that could be utilized in further studies of bid-ask spreads on the NYSE.  $^{8}$ 

#### FOOTNOTES

- 1 NYSE, Fact Book, 1976.
- 2 NYSE, Special Membership Bulletin, May 27, 1976.
- 3 As of the end of 1976, one specialist unit had competition from another unit on the floor. There is of course some competition from market makers located away from the NYSE, and this competition is scheduled to increase with the development of a national quotation system.
- 4 See H. STOLL [7] for a more detailed analysis.
- 5 If names of units given in <u>Stocks and Specialists</u> were not identical they were treated as separate units. It appears from the names only that some specialist firms participated in more than one unit.
- 6 The daily price and returns files of the Center for Research in Securities Prices of the University of CHICAGO was utilized.
- 7 This is the so called "Lawrence Fisher Effect" [2] which has recently been investigated in some detail by Schwartz and Whitcomb [6].
- 8 For studies of bid-ask spreads on the NYSE see Demsetz [1] , Tinic [8] and Hamilton [3]. Only Tinic utilizes a variable the number of specialist stocks which in part reflects diversification but also reflects diseconomics of trading a large portfolio.

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- 8. S. Tinic, "The Economics of Liquidity Services," Quarterly Journal of Economics, 86 (February 1972).

	NUMBER		v.	EL TAL ISI	SPECIALIST UNIT PORTFOLIO CHARACTERISTICS	OLIO CHAR	ACTERISTI	CS		RANDOM	RANDOM PORTFOLIO	CHARACTERISTICS	IISII
	OF COMMON		EQUAL	EQUAL WEIGHTS			MARKET VAI	VALUE WEIGHTS	S	EQUAL	EQUAL WEIGHTS	MARKET VALUE WEIGHTS	STH
SPECIALIST UNIT	CVOOTE	Ď K	D <sub>k</sub>	σ <sup>2</sup> (R <sub>k</sub> ) × 10,000	× 10,000	<sub>గ్</sub> రా	بر ت	σ <sup>2</sup> (R <sub>K</sub> ) × 10,000	R <sub>k</sub> * 10,000	D K	σ <sup>2</sup> (R <sub>k</sub> ) × 10,000	×°	σ <sup>2</sup> (R <sub>k</sub> ) × 10,000
ADLER COLEMAN	43	. 970	.143	58	- 12.83	000.1	.155	1.71	- 10.85	.121	1.58	.213	1.7
BEAR STEARNS, JOHN MUIR	70 77	. 772	. 299	1.23	- 18.69	.803	.351	1.43	- 5.96 - 7.49	.367	2.23 1.89	.443	2.79
BENTON, CORCORAN, LEIB	19	.832	.159	2,13	- 12,49	1.046	.213	2.00	3,89	.174	1.58	. 245	1.96
BENTON, TOMPANE	32	1.149	.122	2.17	8.88	.911	.245	1.59	- 9.63	.158	1.77	.316	2.32
CHAUNCEY	œ ü	.754	.327	2.14	- 11.42 - 6.75	1.135 .675	.336	2.80	- 3.38	.328	1.84	388	2.37
COLIN HOCHSTIN	1 7 8	1.033	.243	2.03	- 11.72	.974	.502	2.74	- 10.51	. 264	1.82	335	3.02
CONKLIN, CAHILL	33	1.106	.128	2.02	- 8.40 - 14.51	1.109	.370	2 3/4	- 6.21	.252	1.85	.344	2.28
CREEM & CREEM	15	.901 775	.294	1.65	9.83	.930	.377	2.00	3.04	.261	1.74	.324	8.) m
	17	.853	.401	1.75	10.08	.878	.492	2.19	- 7.75	.282	1.84	.357	2.28
HAUPT, ANDREWS, FRAIMAN & HU. EINHORN	12 5	.976	. 223	1.76	- 11.42	1.082	.322	2.49	8.61	.291	2.03	.347	N N
ERNST & CO, WARE & KEELIPS	3 2	. 962	.240	1.76	- 10.22	.911	.428	2.09	3.05	.229	2.17	.398	2.63
FARRELL & Co, ROTHENBURG	27	.958	. 198	1.65	- 6.69 - 12.90	.944	. 286	2.67	- 10.06 - 7.80	.184	1.54	. 265	2.00
FOUTER & CO, F.L. SALOMON	15	.991	.300	2.02	9.03	1.162	.282	2.71	- 6.64	.300	1.79	. 362	<b>3</b> K
FRANK, TRAVERS & HUME	38 20	1,110	.152	3 37	- 15.29	1.118	.279	2.49	- 11.21	.138	1.53	. 228	12 1
	26	.970	. 248	1.80	- 9.91 - 14.29	1.078	.210	2.12	- 5.93	.241	1.75	338	2, 15
	-4	1.063	.221	2.09	- 8.81	881.1	.291	2.87	11.00	.305	. 64 	368	2.14
GENGLER BROS	20 9	.912	259	1.62	6.91	1.056	.265	2.19	- 2.19	.228	1.69	.310	2.28
	20	.952	.214	1.66	- 10.37	.995	.492	2.81	. 91	.385	2.34	.444	3.10
	61	.927	229	1.68	9.90	1.025	.273	2.48		.219	1.79	.317	2.08
GOLDSTEIN, LIEF, WERLE		.907	.796	5.81	- 6.79	.861	. 668	3.22	1.36	.596	1.6/	.317	- 2
HIRSHON BOTH	243	. 965	.173	1.62	- 9.44	.896	. 267	1.58	- 10.93	.121	1.57	. 193	- +
JACOBSEN BENJAMIN & SONS	20	.975	. 206	1.86	9.10	1.053	.240	2.10	1.27	.213	1.75	.312	2.26
KINGSLEY, BOYCE, SOUTHWOOD	36	.912	.134	1.38	- 11.29	1.139	.212		- 5.24	153	  	320	- 2
LASKER, SIONE & STERN	ب ا	1.843	. 255	1.37	7.93	.660	.473			252	1.72	.350	2 -
LAURO & Co., M.J. COHEN	- 0	1.013	. 237	1.94	- 12.78	1.090	189			.166	1.68	.256	2
LENARI, METGH	22	.857	. 262		- 10.32	.916	.304	1.74	- 5.22 - 4.72	257	1.84		٠,٧
CIEF, KERLE	35	.802	.170			.803	. 235			.143	1.62	.226	- 19 19 19 19
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	NUMBER			PECIALIST	SPECIALIST UNIT PORTFOLIO CHARACTERI	OLIO CHAR	ACTERISTICS	cs		RANDOM	RANDOM PORTFOLIO	CHARACTERISTIC
	OF		EQUAL	EQUAL WEIGHTS			MARKET VA	VALUE WEIGHTS	S	EQUAL	EQUAL WEIGHTS	MARKET VALUE WEIGHTS
SPECIALIST UNIT	STOCKS	π,	ج ت	σ <sup>2</sup> (R <sub>k</sub> ) × 10,000	× 10,000	ř,	D <sub>k</sub>	σ <sup>2</sup> (R <sub>k</sub> ) × 10,000	R. x 10,000	r K	$\sigma^{2}(R_{k})$ x 10,000	D <sub>k</sub>
ADLER COLEMAN ASIEL	43	.970	.143	1.58	2.	1.000	, 155	1.71	- 10.85	.121	1.58	.213
BEAR STEARNS, JOHN MUIR	17	.772	.299	1,23	- 18.69	.803	. 444	2.67	- 5.96	.367	2.23	.443
BEAUCHAMP & CO BENTON, CORCORAN, LETB	29 19	1,114	.159	2,13	- 12.49	1.046	.213	2.00	- 3.89	.174	1.58	.245
MPANE	32	1.149	.122	2.17	- 8.88	. 911	.245	1,45	- 9,63 - 70	.244	1.77	.316
RODI GAVIN	, <del>I</del> 3	1.00	.327	2.14	- 11.42	1.135	.336	2.80	3.38	.328	1.84	388
DE CORDOVA, COOPER, KAVANEWS	<u>-</u> ∞ α	1.033	24.3	1.49	- 6.75	.675	.583	1.57		.424	2.25	.502
COLIN HOCHSTIN	17	.789	. 286	1.26	- 8-40	786	370	2.74	- 10.51	.264	1.82	.335
COUKLIN, CAHILL	: w	1.106	.128	2.02	- 14.51	1.109	.244	2.34	- 5.69	.162	1.85	344
CREEM & CREEM	15	.775	478	6.0	9.83	.930	.377	2.00	- 3.04	.261	1.74	.324
CUNIFF, LAMM, STOUTENBURGH	17	.853	.401	1.75	- 10.08	.822	.347	1.49	- 4.98	.282	1.84	.357
EINHORN	12	1.174	.187	2.44	8.03	1.082	322	2.49		.291	2.03	.342
ERNST & CO, WARE & KEELIPS	21	.962	.240	1.76	- 10.22	.911	.428	2.09	3.17	. 229	2.17	398
FARRELL & Co, ROTHENBURG	27	.913	. 198	1.65	- 6.69	1.149	.286	2.67	- 10.06	.184	1.54	.265
FOSTER & Co, F.L. SALOMON	15	.991	.300	2.02	9.03	1.162	287	2.71	6.60	300	1./1	. 285
FRANK, TRAVERS & HINTE	3 %	.943	. 152	1.51	- 15.29	1.118	.279	2.49	- 11.21	.138	1.53	. 228
	26	.970	. 248	1.80	- 14.29	1.078	.210	2.12		.241	1.75	.338
	3 -	1.063	. 221	2.09	8.81	1.188	.291	2.87	- 11.00	.305	1.83	368
GENGLER BROS	9 0	.871	. 359	1.62	- 6.9!	1.056	.265	2.19	,	. 228	1.69	.310
GIANNI	20	.952	. 214	1.66	- 14.98	1.091	.310	2.48	9  - 9.78	2000	2.34	.444
GOLDSTEIN, LIEF, WERLE	- 4	.927	.229	1.60	9.90	1.025	273	2.08	- 5.45	.235	1.67	.317
SS	ΰ,	. 965	.173	1.62	0.79	861	.668	3.22	- 1.36	.596	3 <b>.3</b> 8	.649
	21	1.012	. 206	1.86	6.10	1.053	.240	2.10	- 10.93 1 27	.121	57	.193
VINCSUSY BONGE CONTENTON	20	.975	.284	1.91	9.38	1.014	.437	2.63	6.06	. 237	.~ . 	390
LA BRANCHE & Co	10	.912	255	1.38	- 11.29	1.139	.212	2.37	- 5.24	153	1.5	.254
LASKER, SIONE & SIERN		1.000	117	3 5/		.660	.473	1.19	.35	.252	1.72	350
Co, M.J.	19	1.013	. 237	1.94	- 12.78	.880	.389	2.11	- 7.76 - 5.73	.166	1.68	.256
LENART, MCHUGE	22	.857	.262	1.43		.916	.304	74		3 /		
LIEF, WERLE	35	.802	.170	1.12	- 8.52	.803	.235	1.74	- 4.72 - 1.93	.214	1.79	.317
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TABLE 1 (continued)

NUMBER SPECIALIST  OF EQUAL WEIGHTS  COMMON	SPECIALIST UNIT STOCKS $b_k$ $D_k$ $\sigma^2(R_k)$ $\times$ 10,000	MARCUS 21 1.055 .188 1.97 MARCUS SCHLOSS 12 .863 .308 1.55 MEEHAN 17 1.273 .141 2.71	, BREGMAN 21 1.23 .204 1, JONES, TEMPLETON 18 .949 .293 7, MARKETILES & SMITTH 23 .064	15 1.124 .216	PHELAN, SILVER 14 .745 .333 1.20 PHELAN, SILVER 40 1.041 .117 1.77 PICOLI	1.006	. 246	RCELL, GRAHAM	MACK 18 .845 .285	EDY 10 1.236 .311 18 .990 .285		.993 .173	EY, THOMAS 23 1.074 .201 .278	, ROMANO 24 .941 .198 .16 .727 .245	SMITH 13 .852 .303	1ATION 13.1 .122 .111	N 19 .970 .237	.796	4 . 727 . 062
T UNIT	(k) R (000 x 10,000	1 1 1	1 1 1	2 - 6.98	1 1	7 - 6.93	<u></u>	3 - 9.13	6  - 8.92	1 !	1 1	1 1	<del>-                                    </del>	1 1	<u> </u>	7 - 10.75 5 2.98		· <del></del>	<del></del>
PORTFOLIO CHAR	×a,	1.085 1.053	1.365	.914	.950	1.095	776	1.144	.989	1.544	1.042	.920	1.002	.938	1.044	.164	1.015	1.544	. 636
CHARACTERISTICS MARKET VALU	א D	,233 ,260	.19)	.323	.505	.167	.327	.228	.075	.386	.368	.231	.356	.338	.340	.321	.320	.668	.075
CS WEIGHTS	σ <sup>2</sup> (R <sub>k</sub> ) x 10,000	2.21 2.16 3.26	3.26 3.62 1.91	1.57	2.11	2.07	2.66	2.44 3.82	2.01	3,35	2.48	2.81	2.06	1.92	2.38	2.27	2.12	5.59	1.18
· σ	× 10,000			- 1.53 - 3.77		7.69			- 4.18				- 10.56	- 4.73 - 6.77	- 5.68	5.36 3.69	- 5.24		- 11.34
RANDOM EQUAL	, J	.211	.277 .235 .280	.159	.139	.094	.398	.317	.255	.384	.242	.183	.321	.206	.301	.249	.252	.596	.061
PORTFOLIO WEIGHTS	σ <sup>2</sup> (R <sub>k</sub> ) × 10,000	1.72	1.71	1.83					1.44	2.10	1.72	1.75	1.95	1.69	2.10	1.82			1.44
CHARACTERISTICS MARKET VALUE WEIGHTS	D <sub>K</sub>	.276 .391	.347 .326 .360		.382			395	901.	.438	.307	200	3.00	.299	ان کا ان کا	.324			
ARACTERISTICS MARKET VALUE WEIGHTS	σ <sup>2</sup> (R <sub>k</sub> ) x 10,00	2.08	2.14 2.08 2.17	1.96 2.29	2.44 1.95	1.73	2.56	2.00	1.94	2.89	2.00	2.04	2.52	2.23	2.49	2,26	2:24	4.00	

SELECTED RANK VALUES OF DIVERSIFIABLE RISK AS A PERCENTAGE OF TOTAL RISK FOR PORTFOLIOS OF 67 NYSE SPECIALIST UNITS

TABLE 2

	Market Value Weights	ue Weights	Equal Weights	eights
Number of Units with Lower Value	% Diversifiable Risk D <sub>k</sub> x100	Number of Stocks in Portfolio	% Diversifiable Risk D <sub>k</sub> ×100	Number of Stocks in Portfolio
0	7.54	93	6.24	93
Ø	19.73	17	12.24	32
13	23.50	35	16.05	33
20	25.86	21	18.86	17
27	28.63	27	21.39	20
34	32.19	16	24.01	21
41	33.74	12	26.19	22
48	36.76	19	29.39	18
55	41.34	19	31.13	10
62	50.17	18	40.10	17
66	66.82	4	79.60	4



